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Chinese Investment in Mexico: Trade Wars, Nearshoring, and Place-Based Policies

Agustina Giraudy

giraudy@american.edu

American University and School of Social Sciences and Government, Tecnológico de Monterrey

Ernesto Stein

ernestos@tec.mx

School of Government and Public Transformation, Tecnológico de Monterrey

Francisco Urdinez

furdinez@uc.cl

Pontificia Universidad Católica de Chile

Victor Zuluaga

victorzg@banxico.org.mx

Research Department, Banco de México

School of Government and Public Transformation

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Agustina Giraudy † Ernesto Stein ‡ Francisco Urdinez § Victor Zuluaga ¶

Abstract

This paper investigates the impact of the first Trump administration's (2016-2022) US-China Trade War on the sectoral composition and geographic allocation of Chinese foreign direct investment (FDI) in Mexico. Leveraging a novel dataset of Chinese investment projects (2001-2024) and exploiting product-level variation in US tariff exposure, we implement a difference-in-differences design to identify causal effects. The analysis reveals three key findings. First, Chinese firms responded to increased US tariffs by relocating production to Mexico (i.e., nearshoring), with sectors more exposed to the Trade War (that is, receiving larger tariff hikes) having significantly higher Chinese FDI inflows. Second, these effects emerge with a lag of approximately three to five years following tariff imposition. Third, place-based policies significantly influenced the geography of Chinese investment: Mexico's Zona Libre de la Frontera Norte program altered the relative attractiveness for Chinese FDI of the affected regions compared to others. The findings highlight how global trade disputes interact with place-based policies to shape investment patterns, offering lessons for developing economies seeking to attract nearshoring FDI while balancing employment and regional development objectives.

Keywords: Foreign Direct Investment, Trade War, Nearshoring, Mexico, China, Placebased Policy

JEL Classification: F13, F14, F21, F23, O19, P33

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[†]American University, Washington DC, United States of America, and Instituto Tecnológico de Monterrey, Mexico City, Mexico. Email: giraudy@american.edu.

[‡]Instituto Tecnológico de Monterrey, Mexico City, Mexico. Email: ernestos@tec.mx.

[§]Pontificia Universidad Católica de Chile, Santiago de Chile, Chile. Email: furdinez@uc.cl.

[¶]Research Department, Banco de México, Mexico City, Mexico. Email: victorzg@banxico.org.mx.

1 Introduction

The intensification of US–China trade tensions since 2018 has substantially altered patterns of global investment and production. The United States imposed tariffs of up to 25 percent on roughly \$250 billion of Chinese imports, creating powerful incentives for Chinese firms to shift production toward third countries with preferential access to US markets. These elevated tariff levels have remained in effect through 2019–2024, sustaining their influence on cross-border investment flows and the reorganization of global supply chains.

This paper investigates the impact of the first Trump administration's (2016-2022) US-China Trade War on the sectoral composition and allocation of Chinese foreign direct investment (FDI) in Mexico. Mexico, with its geographic proximity to the United States and membership in the USMCA trade agreement, emerged as a prime destination for nearshoring. We focus on two complementary dimensions of this investment: the sectoral reallocation of investment toward products most affected by US tariffs, and the spatial distribution of these investments across Mexican states in response to domestic policy interventions. While these interventions were undertaken at the federal level, we leverage the fact that they had differential impact in different subnational jurisdictions.

Building on insights from multiple strands of scholarship, we advance three hypotheses. First, drawing on the literature on firms' location decisions (Helpman, Melitz, and Yeaple 2004; Tintelnot 2017), we hypothesize that Chinese firms increased their export-platform FDI in Mexico to retain access to the US market while limiting the effect of trade barriers. Specifically, we expect Chinese firms to increase their investments in Mexico following the US-China Trade War, with larger increases in products more exposed to the increases in US tariffs to Chinese goods (Hypothesis 1: Nearshoring). Second, given the complexity of FDI decisions and policy uncertainty during the first Trump administration, we expect investment responses to emerge with significant temporal lags, materializing primarily after policy permanence became clear under the Biden administration (Hypothesis 2: Delayed

Response Pattern). Third, considering the role of labor-cost differentials for multinational location decisions (Konings and Murphy 2006; Muendler and Becker 2010), we expect that Chinese investments will avoid Mexican regions where domestic policy changes raised labor costs, even when combined with tax reductions (Hypothesis 3: Spatial Selectivity).

We test these hypotheses using a difference-in-differences research design that exploits product-level variation in US tariffs imposed during the Trade War, in combination with a novel dataset of Chinese investment projects in Mexico spanning the period 2001–2024. Due to the limitations of official data on Chinese FDI for the Mexican economy, this type of studies have been impossible to date. Mexico is a particularly strong case for testing our hypotheses because it combines structural, institutional, and policy conditions that closely align with the mechanisms we seek to evaluate. Its geographic proximity to the United States and preferential access to its market under the USMCA make it one of the main beneficiaries of trade diversion effects stemming from the US-China Trade War, creating strong incentives for firms to relocate production. At the same time, Mexico's federal structure produces meaningful subnational heterogeneity in tax regimes, labor regulations, and investment promotion strategies, offering an ideal setting to observe how global shocks interact with domestic institutional diversity. Finally, the introduction of territorially targeted policies, such as the Zona Libre de la Frontera Norte, provides a quasi-experimental environment to assess how place-based incentives condition the spatial distribution of foreign direct investment. These features make Mexico uniquely suited for empirically testing our propositions.

Our empirical analysis yields three key findings. First, products facing larger US tariffs experienced significantly larger increases in Chinese FDI flows, with a one percentage-point increase in US tariffs leading to about 5 percent higher Chinese investment amounts at the project and industry level. Second, these effects operated through both extensive and intensive margins—increases in both the number of investment projects and average amounts invested—with effects materializing three to five years after initial tariff implementation.

Third, while most Chinese investments concentrated in US-border states, Mexico's Zona Libre de la Frontera Norte (ZLFN) program actually reduced Chinese investment inflows, as substantial minimum wage increases outweighed value-added and income taxes reductions for export-oriented manufacturers. These findings offer robust empirical support for the nearshoring hypothesis while illuminating how domestic policy choices shape the distribution of benefits from trade-induced investment diversion.

Our results contribute to three bodies of literature: trade war effects on global production networks, the determinants of export-platform FDI location decisions, and the role of place-based policies on the geography of foreign direct investments. The temporal patterns we document have important implications for understanding investment responses to trade policy uncertainty. Our findings illustrate that while trade policy changes can shift investment flows to third countries, certain place-based policies may inadvertently alter the attractiveness of regions to export-oriented investments. These contrasting outcomes provide valuable lessons for host countries designing policies to effectively capture benefits from global production reorganization and recent geopolitical tensions.

The remainder of the paper is organized as follows. The next section places our paper in the literature and describes the theoretical framework behind our empirical work. Section 3 introduces our data and shows the recent evolution of Chinese investment in Mexico at the national and subnational levels. Section 4 describes the empirical strategy and the results using variation in product-level tariffs during the US-China Trade War. Section 5 presents our results for the ZLFN program. Section 6 discusses the policy implications of our work and concludes.

2 Theoretical Framework

2.1 Trade Wars and Nearshoring

The relationship between trade policy and FDI has long been central to international economics research. Classic theories suggest that when trade barriers increase between two countries, firms may respond by establishing foreign production facilities to serve markets that become inaccessible through exports (Helpman, Melitz, and Yeaple 2004). This "tariff-jumping" motive for FDI provides the theoretical foundation for understanding investment diversion during trade disputes. In many cases, such diversion takes the form of export platform FDI, where firms invest in a third country not to serve its domestic market, but to use it as a base for exporting to other destinations (Tintelnot 2017)—a strategy that can also reflect nearshoring when the destination country is geographically proximate to final consumers. Building on this framework, Chinese firms facing higher US tariffs have incentives to establish Mexican operations that can serve the US market while avoiding the negative effects of tariffs (Huang et al. 2023; Jiao et al. 2024) and reducing policy uncertainty (Benguria et al. 2022). We expect the magnitude of this incentive to vary by tariff intensity, creating testable predictions about sectoral investment patterns.

More broadly, the concept of nearshoring—the relocation of production activities to countries that are geographically proximate to final markets—has gained prominence as firms seek to reduce supply chain risks, improve responsiveness to demand fluctuations, and navigate geopolitical changes. These trends have been accelerated by the COVID-19 pandemic and subsequent supply chain disruptions, with many firms reassessing the costs and benefits of geographically dispersed production networks. Recent research has documented various drivers

¹To see how free trade agreements such as the USMCA can affect the attractiveness of a member country like Mexico as a location of FDI from source countries outside the free trade area, see Levy-Yeyati, Stein, and Daude (2003).

²Relatedly, Vortherms and Zhang (2024) document larger exit rates for US firms in China in industries more exposed to the Trade War.

of nearshoring decisions, including rising labor costs in traditional manufacturing centers, increased transportation costs, and growing concerns about supply chain resilience (Gereffi 2014; Gereffi 2025). In practice, nearshoring often leads to export platform FDI, particularly when a proximate country like Mexico offers favorable trade access to key markets.

Changes in tariffs can also lead to a shift in trade patterns in third markets, something that has been documented in the context of the first Trump administration's US-China Trade War. For instance, Fajgelbaum and Khandelwal (2022), Fajgelbaum, Goldberg, Khandelwal, et al. (2024), and Freund et al. (2024) demonstrate significant trade diversion effects, with US imports shifting toward countries like Mexico and Vietnam after the imposition of tariffs to Chinese goods. However, the foreign investment dimension of this reallocation has received less attention, particularly at the subnational level where policy variation can shape the geography of foreign investments.^{3,4}

Drawing in this body of literature, our first hypothesis is that Chinese firms strategically increased their foreign direct investment in Mexico following the US-China Trade War, with larger increases in sectors facing higher US tariffs (**Hypothesis 1: Nearshoring**).⁵

2.2 Temporal Dynamics of FDI

Another important dimension of the analysis concerns the temporal dynamics of investment responses to policy changes and uncertainty. The permanence of US tariffs remained unclear during the first Trump administration, creating uncertainty about whether trade measures

³In a related paper, Schulze and Xin (2025) show that the US-China Trade War led to increased Chinese FDI in Vietnam. In addition to our focus on Mexico, our paper is distinct to them in two important dimensions: First, our empirical work leverages variation in tariffs at the product level and, therefore, we identify effects by comparing products not countries, which improves the internal validity of our estimates. Second, we also leverage geographic variation, in particular, the implementation of the ZLFN, to study the role of place-based policies in shaping the spatial distribution of Chinese FDI.

⁴In terms of domestic investments, Caldara et al. (2020) shows that the increase in trade policy uncertainty during the US-China Trade War led to a decrease in corporate investments in the United States. Similar results have been found for Chinese firms in the energy sector (Li et al. 2023).

⁵Historical precedent supports Mexico's role as a nearshoring destination. Paus and Gallagher (2008) document how foreign investors have previously used Mexico as a manufacturing platform, something that partially changed after China's entry into the WTO (Gallagher, Moreno-Brid, and Porzecanski 2008). An important example of these dynamics can be seen in the automotive sector, including more recently in electric vehicles (Crossa and Ebner 2020; Martinez and Terrazas-Santamaria 2024).

would persist beyond the 2020 presidential election. The Biden administration's decision in January 2021 to maintain Trump-era trade policy toward Chinese goods provided crucial signal about policy permanence, reducing uncertainty for potential investors.

Economic theory suggests that investment decisions by firms are highly sensitive to uncertainty because of their sunk costs and irreversibility (Dixit and Pindyck [1994] Caballero, Engel, and Haltiwanger [1995]). FDI projects, in particular, involve lengthy processes of site selection, regulatory approvals, contract negotiations, and supply-chain reorganization, often requiring several years to materialize, which can reduce the responsiveness of investors to demand shocks, including changes in trade policy (Bloom, Bond, and Van Reenen [2007]). Political cycles amplify this timing: firms frequently delay commitments during periods of political uncertainty, such as during the 2020 US presidential election, preferring to wait until policy trajectories are clearer (Gulen and Ion [2016]).

These factors help explain why investment responses to the US-China Trade War were unlikely to be immediate. Instead, we expect the most pronounced increases in Chinese FDI to Mexico to appear a few years after the initial tariff changes, once policy continuity became evident and adjustments were completed. Evidence from other contexts, such as the delay in the adjustments made by European firms after the 2018-2019 trade tensions and the supply-chain disruptions during the COVID-19 pandemic supports this interpretation (EIB 2025).

Our second hypothesis then addresses these temporal dynamics: Given the complexity of FDI decisions, policy uncertainty about tariff permanence following the first Trump administration, and typical investment lumpiness, Chinese FDI responses should exhibit delayed effects, with the strongest investment flows materializing after president Biden's inauguration. (Hypothesis 2: Delayed Response Pattern).

2.3 Subnational Policy Variation and FDI Location

A growing body of research examines how subnational policy variation influences FDI location decisions within countries (Samford and Gómez 2014; Garriga 2022; Giraudy, Urdinez, and Freites 2024). This literature builds on the insight that multinational corporations face location choices not just across countries, but also among regions within countries that may offer different policy incentives, factor costs, and institutional environments.

For instance, Garriga's (2022) analysis of partisan effects on FDI across Mexican states provides important evidence that subnational political characteristics, such as left-leaning state governments, significantly influence foreign investment patterns. Similarly, recent work by Giraudy, Urdinez, and Freites (2024) shows that multilevel partisan alignment between national and subnational governments affects Chinese extractive investments in Latin America, revealing the importance of political coordination across government levels.

Federal systems are generally more likely than unitary systems to exhibit regional variation in policy incentives, factor costs, and institutional environments relevant to investors. This variation stems from the considerable autonomy that subnational governments often wield over key areas such as economic development policy, taxation, and labor regulation (Samford 2022). In addition to this decentralized policymaking authority, national governments in federal countries may also implement territorially targeted policies that affect specific regions (i.e., place-based policies). Together, these dynamics generate substantial interregional heterogeneity, offering foreign investors the opportunity to direct FDI in subnational units where incentives are maximized.

Drawing on this literature, we examine the impact of place-based policies on the spatial distribution of Chinese investment. Our analysis focuses on a specific place-based policy, Mexico's ZLFN program, a regional development initiative launched in January 2019, during the López Obrador administration (2018-2024). The ZLFN was implemented in 43 municipalities along the US-Mexico border and combined three core policy instruments that exerted

divergent effects on production costs. Specifically, the program reduced the value-added tax (VAT) from 16% to 8% and lowered the corporate income tax (ISR) from 30% to 20%. However, the ISR reduction was not automatic; it required firms—particularly new entrants—to apply for and obtain approval from the Ministry of Finance (Hacienda). At the same time, the program mandated significant increases in the minimum wage, substantially reshaping the labor cost structure for manufacturing firms. Together, these measures created a complex policy environment with both incentives and disincentives for investment, offering a valuable setting for analyzing how targeted fiscal and labor policies influence the locational choices of foreign investors, including those from China.

The minimum wage component represents the most dramatic policy change in the ZLFN program. When initially implemented in 2019, the border minimum wage was set at 176 Mexican pesos per day compared to 88 pesos in the rest of Mexico. The minimum wage had been uniform across the country at 80 pesos per day in 2018. However, the most significant changes occurred subsequently: by 2024, the border minimum wage reached 420 pesos per day compared to 278 pesos in the rest of the country, creating a 51% premium. The 425% increase in minimum wage for the border municipalities between 2018 and 2024 contrasts with a rate of inflation of just 33.9% for the same period, resulting in very significant increases in the minimum wage in real terms. These substantially higher wages, justified by the higher costs of living in border municipalities, created a significant cost disadvantage for labor-intensive manufacturing operations.

The geographic coverage of the ZLFN creates important variation even within border states (Calderón et al. 2023). Major industrial cities like Tijuana, in Baja California, and Ciudad Juárez, in Chihuahua, are included in the ZLFN, while other significant manufacturing centers like Monterrey in Nuevo León and Saltillo in Coahuila are located in border states but outside the ZLFN municipalities. This geographic variation is crucial for understanding nearshoring location decisions: firms seeking proximity to the US market for logistical advantages might strategically locate in non-ZLFN municipalities within border states to

capture transportation benefits while avoiding higher labor costs.

Even though the theoretical prediction for the ZLFN's effect on Chinese investment is ambiguous, we conjecture this policy reduced export-oriented foreign investments. This prediction emerges from the asymmetric impact of the policy components on different types of firms. The tax benefits primarily favor domestic-market oriented firms since VAT reductions are irrelevant for exporters, who do not pay VAT on exports and typically do not pay VAT on inputs due to other programs (for example, the IMMEX). While the ISR reduction provides some benefit to firms, this advantage must be weighed against substantially higher labor costs that affect all operations regardless of market orientation, but that have been shown are of particular importance for multinational in emerging economies (Konings and Murphy 2006; Muendler and Becker 2010). For Chinese firms pursuing nearshoring strategies, the cost-benefit calculus is particularly unfavorable. Since these firms are primarily interested in Mexico as an export platform to serve the US market, the VAT reduction is mostly irrelevant. Similarly, there was uncertainty about the income tax's permanence. In contrast, the minimum wage increases represent substantial cost increases for labor-intensive manufacturing operations, particularly as wage levels escalated through 2024 making the minimum wage binding.

Thus, our third hypothesis concerns the geographic and sectoral distribution of investment flows: Chinese firms strategically concentrated their nearshoring investments to avoid Mexican border municipalities where the ZLFN program's minimum wage increases outweighed tax reduction benefits for export-oriented manufacturers (**Hypothesis 3: Spatial Selectivity**).

3 Data for Analysis and Stylized Facts

3.1 Data

To test these hypotheses, we developed a novel dataset that compiles and systematizes all identified Chinese FDI projects in Mexico between 2001 and 2024. Our dataset adopts a project-level approach that permits closer examination of the spatial, economic, and political implications of Chinese FDI. The dataset follows the OECD definition of foreign direct investment, which requires a minimum of 10% equity ownership by the foreign investor. This threshold serves as a proxy for sustained managerial influence and meaningful investment presence. By grounding case identification in an internationally recognized standard, the dataset avoids conflating FDI with other forms of economic engagement such as portfolio investment, procurement contracts, or supplier agreements.

Only projects that were implemented, rather than merely announced, are included. This choice reflects an emphasis on realized capital flows and on-the-ground economic presence rather than investment intentions or strategic signaling. The dataset captures three primary entry modes: greenfield investments, acquisitions, and joint ventures where Chinese investors hold at least 10% of the capital. Greenfield investments constitute the dominant mode, representing 95% of identified Chinese FDI projects in Mexico (172 projects), followed by acquisitions (9 projects). 9 projects are joint ventures, which are not mutually exclusive with the other types of investment. This distribution reflects the empirical patterns of Chinese corporate strategy in Mexico rather than a methodological restriction.

The investment identification and verification process relied on a combination of primary and secondary sources. Each project was cross-checked using at least three independent and credible sources before being included in the dataset. For every verified investment, coders documented the investor's corporate origin, the amount invested (when reliably available), the year the project was implemented, the project's industrial purpose or subsector, and its

location in Mexico.

Finally, the coding rules were designed to ensure comparability, transparency, and conservative estimation. When reported amounts differed across sources, the lowest amount was recorded to avoid inflating project size. Investment value was imputed to the first year of project implementation, even when disbursement extended over multiple years, in order to standardize temporal coding and facilitate longitudinal analysis. Projects with unverifiable amounts were retained but coded as missing in the "amount" field rather than excluded altogether. Likewise, investments originally launched by Chinese companies but later acquired by non-Chinese firms remain in the dataset as evidence of Chinese entry into the Mexican market and its subsequent evolution.

Total Chinese investments in our dataset amount to \$13.6 billion across 181 projects, with investment amounts unavailable for 19 projects due to missing or unverified data. These figures position our dataset between estimates from other rigorous academic sources. The Monitor of Chinese OFDI in Latin America and the Caribbean (MOFLAC) (Dussel Peters 2025) reports \$22.47 billion for Mexico during 2000-2023 across 166 transactions, while the China Global Investment Tracker (CGIT) (Scissors 2024) reports \$6.06 billion for Mexico during 2000-2024 across 26 transactions. Key methodological differences explain these differences. Unlike MOFLAC or CGIT, our dataset excludes infrastructure projects where Chinese corporations win tenders to execute public works, instead focusing on FDI in productive capacity. Additionally, we include projects smaller than \$100 million—unlike CGIT—which turns out to be crucial given that the average investment size in our dataset is \$83.9 million per project, reflecting the prevalence of small- and medium-scale investments in Chinese FDI to Mexico.

Our figures diverge dramatically from official statistics: Mexico's Secretaría de Economía reports only \$2.96 billion in accumulated Chinese FDI since 1999 (Secretaría de Economía 2024), while China's Statistical Bulletin of Outward FDI documents merely \$1.3 billion in stock accumulated through 2021 (MOFCOM 2022). These discrepancies are particularly

stark for geographic distribution: while Mexico's Secretaría de Economía identifies Mexico City as receiving 34.9% of Chinese FDI, our dataset shows Mexico City accounts for only 2.5% of realized investments, with the bulk concentrated instead in manufacturing hubs like Nuevo León (24.9%), Coahuila (13.8%), and Guanajuato (10.7%)—reflecting the operational reality of Chinese manufacturing investments versus the corporate addresses captured by official statistics. Other sources have found evidence of underreporting of Chinese investments for political reasons (Federal Reserve Bank of Dallas 2025; Rhodium Group 2023), as well as systematic triangulation of Chinese investments through third countries that official statistics fail to capture (ECLAC 2025; Urdinez and Myers 2025).

The methodological superiority of academic tracking systems like ours lies in identifying investments by their operational reality rather than legal financial structure. While official statistics follow IMF Balance of Payments standards that record the immediate investor nationality (IMF 2022), our methodology tracks the ultimate beneficial owner and final operational destination, consistent with OECD recommendations for excluding Special Purpose Entities (OECD 2023). This distinction is crucial for Chinese investments, where between 50-90% of global flows are channeled through offshore financial centers before reaching their operational destinations (Scissors 2024; Rhodium Group 2023).

We can further illustrate this methodological distinction with prominent cases from our database: First, Hofusan Industrial Park (Nuevo León): While officially registered through a Singaporean entity, this \$300 million development is operationally a joint venture between China's Holley Group and Mexico's Fusan Investment (Prodensa 2024). Second, Giant Motors Latinoamérica (Hidalgo), structured as 50% financed by Mexico's Grupo Inbursa with the remaining investment coming through a Cayman Islands structure controlled by China's FAW Group, this arrangement allows production of Chinese JAC vehicles in Mexico without appearing as Chinese investment in official records (El Universal 2021). Third, CNOOC in deep-water petroleum exploration: Operating through two Netherlands-registered companies (CNOOC Netherlands B.V.), this Chinese state-owned enterprise maintained over 7 billion

pesos in Mexican bank accounts while exploring Gulf of Mexico fields, yet appeared in official statistics as Dutch investment (Damgaard, Elkjaer, and Johannesen 2024; Forvis Mazars 2013).

The 2024 BBVA-AMPIP Industrial Parks Survey provides compelling field validation of our approach: it documents 34 new Chinese companies physically establishing operations in Mexican industrial parks during 2023 alone, with projected capital requirements of \$510-1,020 million based on typical manufacturing project scales (BBVA Research & AMPIP 2024). This operational evidence far exceeds the \$232 million officially reported by China for all of 2021, confirming that our methodology captures the economic reality that official statistics fragment across multiple jurisdictions due to complex corporate structures.

Importantly, our dataset also includes detailed information on the municipality and main product (6-digit HS code) for each project, enabling precise analysis of both geographic and sectoral patterns. For instance, we can track Lenovo's strategic expansion across Mexico's technology manufacturing corridor: their 2009 investment established a computer assembly plant in Monterrey producing automatic data processing machines (HS code 847130), while their 2018 investment expanded capacity with five new production lines at their Apodata plant in the same state, also focusing on computer manufacturing equipment (HS code 847130), demonstrating intensification within a regional cluster. Similarly, we can observe Yanfeng Global Automotive's strategic expansion across Mexico's automotive corridor: their 2018 investment established a technical center in Querétaro municipality producing vehicle body parts and accessories (HS code 870829), while their 2023 investment created a manufacturing plant in Ciénega de Flores, Nuevo León, shifting to seating components for motor vehicles (HS code 940120), demonstrating both geographic expansion and product line diversification within the automotive value chain. These examples illustrate how the granular municipality and HS code data allow us to distinguish between different patterns of Chinese FDI—from cluster intensification strategies to cross-regional value chain expansion.

3.2 Stylized Facts

As seen in Figure 1 the temporal distribution of Chinese FDI in our data reveals a clear acceleration of investments during and after the US-China Trade War period, with the most dramatic surge occurring in 2023-2024 when 59 projects (32.5% of all Chinese investments) were implemented, representing approximately \$5.7 billion in investment value. This delayed response pattern supports our prediction about temporal lags in investment decisions. Projects classified as potentially benefiting from nearshoring increased dramatically from a limited presence in the pre-2019 period to constituting the vast majority of investments in 2019-2024, with manufacturing-sector projects representing \$9.9 billion or 73% of total documented investment (see Figure 2, panel (a)). We also find that most of the increase in Chinese investments during the end of our study period took place across diverse geographic locations, with significant growth in the northern manufacturing corridor (see Figure 2, panel (b)).

Figure 3 illustrates the geographic concentration of Chinese investments, revealing strategic location decisions that align with proximity to US markets. The northern Mexican states of Baja California, Chihuahua, Coahuila, and Nuevo León concentrate 38% of Chinese projects (69 of 181), with Nuevo León leading with 33 total projects representing \$3.3 billion in investment (24.2% of total documented value), followed by Coahuila with 22 projects worth \$1.8 billion (13.2%). This northern concentration becomes even more pronounced in the post-2018 period, with northern states accounting for 48% of projects (52 of 108 projects from 2019-2024) compared to 34% in the pre-trade war period (25 of 73 projects before 2018) (see Figure 4).

⁶We classify as near shoring those projects whose main product was targeted by the increase in US tariffs to Chinese goods between 2018 and 2019.

4 Trade War and Chinese Investments

In this section, we study the effect of the first Trump administration's US-China Trade War on investments by Chinese firms in Mexico. Specifically, we leverage plausibly exogenous variation in the US tariffs on imports from China to estimate the changes in Chinese investments in products differentially exposed to the trade dispute, which allows us to test Hypotheses 1 and 2 outlined above. To preview our findings, the Trade War led to an increase in Chinese FDI inflows into Mexico toward the end of our study period, particularly in products that experienced the largest tariff hikes. These effects take place at both the extensive and intensive margin of investments, that is, in both the number of investments and the amount invested.

4.1 Empirical Strategy

Using project-level data (i.e., unique investment instances), we start by estimating the following difference-in-differences (DID) specification:

$$\log(Amount_{ipt}) = \beta Post_t \times \Delta Tariff_p + \gamma \Delta Tariff_p + \delta_t + \varepsilon_{ipt}, \tag{1}$$

where $Amount_{ipt}$ is the amount invested by firm i producing product p in year t, with p being 6-digit HS codes, $Post_t$ is a dummy equals to 1 from year 2019 onward, $\Delta Tariff_p$ is the percent change in the US tariffs applied to Chinese imports of product p as a result of the Trade War, δ_t are time fixed effects (FE), and ε_{ipt} are the unobservables allowed to be correlated at the 4-digit HS level. To compute the change in tariffs at the 6-digit HS level, we use as weights US imports from China within each code in 2016 and 2017, with the import and tariff data obtained from Fajgelbaum, Goldberg, Kennedy, et al. (2019). We also estimate specifications in which we replace $\Delta Tariff_p$ by a dummy equal to one if product p experienced any increase in tariffs during the US-China Trade War. The coefficient of

interest is β , which measures how the average change in the Chinese FDI after the Trade War compares between products more or less exposed to the increase in US tariffs. Notice that equation (1) identifies a relative effect (according to differential tariff exposure) and, therefore, we are not able to recover in this regression the aggregate impact on investments of the Trade War (Caliendo and Parro [2022)).

As standard in DID designs, there are two underlying assumptions behind specification (I): First, the evolution of the Chinese FDI in products facing differential changes in US tariffs would have been similar in the counterfactual without the Trade War, this is, that there are parallel trends for products irrespective of their Trade War tariff hikes (Angrist and Pischke 2008). In the context of this paper, the previous assumption requires that, on the one hand, changes in US trade policy are not endogenous to Chinese firms' investment decisions in Mexico, this is, that US tariffs to imports from China are not set strategically to limit the reallocation of Chinese FDI to third countries like Mexico and, on the other hand, that investments in industries with different exposures to the Trade War had similar trends before the policy dispute. Second, there are no anticipation effects in investments (Roth et al. 2023), which in our case implies that Chinese investors were not able to anticipate the timing and impacts of the Trade War or that any anticipatory behavior did not change investment decisions significantly. Even though both assumptions are plausible, we provide suggestive evidence of no pre-trends below.

Given the reallocations of US imports toward countries like Mexico after the US-China Trade War (Utar, Cebreros, and Torres 2025; Fajgelbaum, Goldberg, Khandelwal, et al. 2024; Freund et al. 2024), our expectation is that the changes in US tariffs targeting goods from China should have led to a rise in Chinese investments inflows into Mexico in product more exposed to the tariff hikes. The rationale is that higher tariffs on US imports from China make it relatively more profitable for Chinese firms to establish subsidiaries in third countries, like Mexico, to serve directly or indirectly the US market. By doing so, these firms can reduce the negative effects of higher tariffs (Huang et al. 2023; Jiao et al. 2024) and

benefit from lower policy uncertainty relative to continued production in China (Benguria et al. 2022), effectively making use of these locations as export platforms (Tintelnot 2017; Flaaen, Hortaçsu, and Tintelnot 2020).

To test for pre-trends and dynamic effects, we also estimate an event-study specification with the following form:

$$\log(Amount_{ipt}) = \sum_{\tau} \beta_{\tau} I[\tau = t] \times \Delta Tariff_p + \gamma \Delta Tariff_p + \delta_t + \varepsilon_{ipt}, \tag{2}$$

where τ is a set of periods before and after the Trade War, in particular, we consider two periods before (2014 or earlier and 2015-2016) and three periods after (2019-2020, 2021-2022, and 2023-2024), with the reference period being the years 2017-2018, one year before the Trade War and during the first rise in US tariffs to imports from China.⁷ $I[\cdot]$ is an indicator function that equals 1 if the condition in the brackets is true. The rest of the variables were defined above. The coefficient of interest in this case is β_{τ} , which measures how the difference in Chinese FDI between period τ and 2017-2018 compares between products with varying exposures to the tariff changes during the Trade War. For periods before the policy change, these coefficients are informative of pre-trends and anticipation effects in Chinese investments in the more exposed products. For periods after the Trade War, these coefficients capture the dynamic effect of the changes in tariffs, which we expect to be important given the lumpiness of investments found in the literature (Caballero, Engel, and Haltiwanger [1995]) and the uncertainty and impacts from the COVID-19 pandemic. Since the trade policy changes implemented during the first Trump administration could have been reversed by the subsequent administration, investors may have also faced uncertainty about the tariffs themselves (see our discussion of Hypothesis 2).

As mentioned earlier, equations (1) and (2) are estimated at the project level. That is,

⁷This choice is also consistent with the fact that the increase in US tariffs was much more significant in 2019: Out of a total of \$566 billion of Chinese imports affected by tariffs, only \$56 billion experienced increases in tariffs of 25 percentage points in 2018, which contrast with about \$500 billion having this increase in tariffs in 2019.

each observation is a Chinese FDI project in Mexico, which implies that we are only using information related to the projects that were actually undertaken.⁸ Thus, we are not able to capture, for example, why investments happen in some sectors and not in others. To better capture compositional effects, we aggregate the data to the 4-digit HS level (i.e., heading level). Thus, we can run regressions where the unit of observation is the heading-year. This allows us to include observations in which there was no Chinese FDI. We are then able to estimate the following two-way FE specification:

$$Y_{ht} = \beta Post_t \times \Delta Tariff_h + \gamma_h + \delta_t + \varepsilon_{ht}, \tag{3}$$

where Y_{ht} is the outcome of interest (number of projects, total amount invested, etc.) in heading h in year t, $\Delta Tariff_h$ is the average change in the tariffs imposed by the US to imports from China during the Trade War for heading h, γ_h are heading FE, and the rest of the variables are the same as before. As before, to compute the tariff change at the heading level, we use US import from China as weights. The coefficient of interest in this case is β , which measures how the average change in the outcome of interest after the Trade War compares between headings heterogeneously exposed to the increases in US tariffs. As established before, (3) implicitly assumes parallel trends between headings heterogeneously exposed to the Trade War and no anticipation effects.⁹

Using the heading-level data, we also estimate an event-study specification similar to (2) as follows:

$$Y_{ht} = \sum_{\tau} \beta_{\tau} I[\tau = t] \times \Delta Tariff_h + \gamma_h + \delta_t + \varepsilon_{ht}, \tag{4}$$

⁸Note that if the Trade War led to an increase in the number of projects (i.e., affected the extensive margin), this may introduce selection bias in our project-level regressions. In particular, the set of observed projects post-treatment may differ systematically in unobserved characteristics from those that would have occurred in the absence of the Trade War (i.e., the counterfactual), affecting our inferences at the project level.

⁹Notice that in (3) we are able to control for heading fixed effects, which allows us to account for unobservable factors fixed over time at the heading level and, therefore, improves the internal validity of the estimates. Including similar fixed effects in (1) and (2) is infeasible given the granularity of the data in those regressions.

with β_{τ} being the effect of the change in tariffs resulting from the Trade War between period τ and 2017-2018. As before, we estimate effects for two periods before (2014 or earlier and 2015-2016) and three periods after the Trade War (2019-2020, 2021-2022, and 2023-2024). Similar to our analysis at the project level, we use the coefficients in (4) for periods before the Trade War to test for pre-trends and anticipations effects. The coefficients for periods after the policy change in (4) allow us to document dynamic effects.

4.2 Results

4.2.1 Project-level results

We start our empirical analysis estimating equation (1) with project-level data. Table 11, Panel A, reports the results for several specifications, each with a more stringent set of controls: column (1) shows the results with only the change in tariffs ($\Delta Tariff_p$) and year FE as controls, columns (2) and (3) add manufacturing and region FE, and columns (4) and (5) add linear trends for manufacturing firms and each region separately. Including FE and linear trends for manufacturing firms and regions controls for systematic differences in the levels and growth of investments along these two dimensions (for example, investments might be larger or growing faster in the manufacturing industry). Conservatively, we cluster the standard errors at the four-digit HS level to allow for correlation in the error terms of closely related products (Abadie et al. 2023).

As seen in Table [1], Panel A, we do not find any average effects of the changes in US tariffs during the Trade War on Chinese FDI in Mexico, with the coefficient of interest being positive and large but imprecisely estimated, a result that is robust to alternative combinations of FE and linear trends. As shown in Appendix Table [A.1], Panel A, replacing the actual change in tariffs by a dummy equal to one if the product experienced any change in tariffs during the US-China Trade War yields similar results.

The results of the difference-in-differences specifications discussed so far, however, may

obscure how Chinese investment responses to US tariffs varied over time during the post-Trade War period, which is likely given that investment decisions tend to be lumpy and irregular. Specifically, investments may respond slowly to trade policy changes, meaning tariff effects on investor behavior could emerge years after the policy shift. This delayed response is especially important to consider since investors during the first Trump administration may have expected subsequent administrations to reverse the tariff increases.

Table [1] Panel B, shows the results of estimating equation (2), in which we allow the effect of US tariffs on Chinese investments to vary flexibly for periods before and after the trade dispute, with the omitted category being the period 2017-2018. We do not find evidence of pre-trends or anticipation effects, with the coefficients for all periods before 2017 being small and statistically insignificant at conventional levels. However, we do find strong and statistically significant effects of the changes in tariffs on Chinese investments in the last period of our analysis (2023-2024), which suggests that investments from China took several years to react to the changes in US trade policy, something consistent with the initial uncertainty regarding the permanence of the tariffs, the lumpiness of investments, and the slow recovery from the COVID-19 pandemic. In terms of magnitude, we find that a one percentage-point increase in US tariffs to Chinese imports led to increases in the average amount invested by Chinese firms in Mexico of around 5 percent, a sizable and economically significant effect suggesting strong reallocation of investments from China toward Mexico in the most exposed industries. The strong and lagged effects of the US tariffs on investments are consistent with our Hypotheses 1 and 2.

Table A.1. Panel B, in the Appendix reports the results of estimating the previous event-study specifications replacing the change in tariffs by a dummy equal to one for products exposed to the US-China Trade War. As can be seen, our main results do not change—and perhaps are even strengthened—in these specifications, with the point estimates for the period 2023-2024 being large and statistically significant, suggesting a large increase in Chinese investments in industries exposed to the Trade War. Our results are also robust to

using simple averages when computing the US tariffs at the 6-digit HS level (see Table A.2 in the Appendix).

Notice that our analysis so far abstracts from the retaliatory tariffs imposed by China to the United States during the Trade War, which may be important for Chinese firms using US inputs when producing goods for the exporting market. Table 2 reports the results of estimating specifications (1) and (2) adding the retaliatory tariffs as controls. As can be seen, although less precisely estimated, we find quantitatively similar effects of the US tariffs on the Chinese investments in Mexico, with large and statistically significant effects for the period 2023-2024. These results are consistent with the differences in the targeting of the US and Chinese retaliatory tariffs, with the former including a wide set of products and the latter being politically targeted (Fetzer and Schwarz 2021), which implies that both sets of tariffs have a low correlation. 10,11

4.2.2 Heading-level results

Specifications (1) and (2) estimate the effect of US tariffs on FDI originated in China but, since all observations correspond to actual investments, we can only look at the intensive margin (what explains the amount invested per project) but not the extensive margin (what explains where there is an investment). To study the joint effect on the number of projects and the amount invested, i.e., the extensive and intensive margins of investments, we estimate equations (3) and (4) using data at the four-digit HS (heading) level. The results can be found in Table (3). In particular, we show the estimates for three dependent variables: A dummy if there was at least one Chinese investment project in that heading-year in column

¹⁰Beijing targeted products related to politically influential sectors, or states that were important to president Trump's coalition. For example, it imposed 25% tariffs on soybeans, a product for which China is the largest market, seeking to hit Midwestern farmers in states such as Iowa, Illinois and Indiana, a key republican constituency with strong lobby capabilities. Similarly, it also targeted US pork products, important for North Carolina, a key swing state, and bourbon, a key product for the state of Kentucky, which at the time was the home state of the Senate Majority Leader.

¹¹In our analysis, we also abstract from the US tariffs to Mexico and the Mexican retaliatory tariffs, which follows from the fact that both set of tariffs were rolled back in May 2019 and that the size, coverage, and impact of these tariffs were much more limited (Utar, Cebreros, and Torres 2025).

(1), the actual number of projects in column (2), and the amount invested transformed with the inverse hyperbolic sine (i.h.s.) function in columns (3), which allows us to include observations without investments in the regressions (Bellemare and Wichman 2020). All specifications control for year and heading FE. We also conservatively control for the Chinese retaliatory tariffs. We cluster the standard errors at the heading level to allow for correlation between observations of the same heading over time.

As seen in Table 3. Panel A, we find positive although small and not statistically significant effects of the changes in US tariffs on the extensive and intensive margin of investments during all the periods after the policy change. Notice also that the fact that the coefficients for both margins are not statistically significant suggests that our results in Tables 1 and 2 using project-level data are not driven by selection bias in the type of project following the Trade War.

The results for the event study using heading-level data can be found in Table 3 Panel B. As before, we report coefficients for periods before and after the Trade War. Consistent with our findings so far, we do not find evidence of pre-trends or anticipation effects, which lends confidence to our identification strategy. For the periods after the Trade War, though, we find that the changes in US tariffs led to statistically significant increases in the extensive and intensive margins in 2023-2024. These effects are also economically significant, particularly for the intensive margin: A one percentage-point increase in the US tariffs to imports from China led to, on average, an 5.4 percent increase in the Chinese investments in Mexico in the period 2023-2024. Again, these patterns are consistent with our Hypotheses regarding the sectoral and dynamic response of investments to tariff changes.

In summary, the empirical analysis in this section suggest that the US-China Trade War

The inverse hyperbolic sine transformation is $\ln[x + (x^2 + 1)^{0.5}]$, where x is the investments in millions of US dollars. Notice that this function is defined when there are no investments (x equals zero) and, therefore, it does not have the limitations of using logs.

¹³As described above, the change in tariffs at the heading level corresponds to an average of the changes at the 10-digit HS level using US imports from China as weights. In Table A.3 in the Appendix, we report the results for specifications (3) and (4) using a simple average of the 10-digit tariffs instead, with the estimates being quantitatively similar to those reported in the main text.

led to a reallocation of the Chinese FDI in Mexico toward products and headings that experienced the largest increases in US tariffs. These findings align with Hypothesis 1, suggesting that Chinese investors are increasingly leveraging Mexico as an export platform to access the US market, motivated by reduced tariff exposure and diminished policy uncertainty. The lagged response is consistent with the lumpiness of investments and the uncertainty in US trade policy (Hypothesis 2). More broadly, the results provide empirical support for the nearshoring narrative, in which geopolitical tensions and rising protectionism incentivize firms to relocate production or investment to countries that offer geographic proximity and preferential market access to the United States (Fajgelbaum and Khandelwal 2022). Mexico's integration into North American supply chains appears to have positioned it as a key destination in this reconfiguration of investment flows.

5 Place-Based Policies and the Location of Chinese FDI

We now turn to the question of where within Mexico the Chinese investments were located. Specifically, we study how domestic policies shape the geography of foreign investment. Our focus is the ZLFN program, a large-scale place-based policy implemented in January 2019 that combined reductions in the VAT and ISR and a substantial increase in the minimum wage in municipalities along the northern border. Our results provide suggestive evidence that municipalities in the ZLFN received comparatively less Chinese FDI. This pattern is consistent with the interpretation that the increase in labor costs, particularly relevant for export-oriented manufacturing, outweighed the benefits from lower consumption and corporate income taxes for foreign investors targeting the US market (Hypothesis 3).

5.1 Empirical Strategy

To estimate the effect of the ZLFN on investments, we start by estimating the following DID specification:

$$Y_{lt} = \beta Post_t \times Border_l + \gamma_l + \delta_{rt} + \varepsilon_{lt}, \tag{5}$$

where l indexes the interaction between state and border, more precisely, we define l such that for non-border states the location is equal to the state and for border states there are two observations, one for border municipalities (as defined in the ZLFN decree) and another for non-border municipalities. 14 Y_{lt} is the outcome of interest (number of projects, total amount invested, etc.) in location l in year t, $Border_l$ is a dummy equal to one for locations along the border with the United States, γ_l and δ_{rt} are location and region-by-time FE, and the rest of the variables are the same as before. We cluster the standard errors at the location level, which is the geographic level at which the policy takes place. The coefficient of interest is β , which measures how the average change in Chinese FDI after the start of the ZLFN program compares between locations along the border with the United States and in the rest of the country. As discussed for the case of the US-China Trade War, β assumes parallel trends between locations along and off the US border and no anticipation effects.

To test for pre-trends and dynamic effects, we also estimate an event-study specification as follows:

$$Y_{lt} = \beta_0 + \sum_{\tau} \beta_{\tau} I[\tau = t] \times Border_l + \gamma_l + \delta_{rt} + \varepsilon_{lt}, \tag{6}$$

where τ is a set of periods before and after the ZLFN program. Given that there are six border states, which limits the precision and power of our estimates, we reduced the number of parameters to estimate by only considering one period before the program (all years before 2015) and two three-year periods after the program (2019-2021 and 2022-2024),

¹⁴Following this definition, there are 37 locations in total, 26 for non-border states and 11 for border states (two observations for the states of Sonora, Chihuahua, Coahuila de Zaragoza, Nuevo León, and Tamaulipas, and one for Baja California where all municipalities are considered to be along the border for the purposes of the ZLFN policy).

with the years between 2016 and 2018 serving as comparison. The rest of the variables were defined above. The coefficient of interest in this case is β_{τ} , which measures how the change in Chinese FDI between τ and 2016-2018 compares for locations along and off the US border. Although less granular and precise than our analysis in the previous section, we can document pre-trends with the periods before the ZLFN program. With the post-treatment periods, we can explore dynamic effects by estimating differential effects for the COVID-19 years (a period during which the minimum wages might not yet have been binding) and the latest years in our study period, where we found effects with the tariff changes during the Trade War.

5.2 Results

Our estimates of the effect of the ZLFN program on Chinese investments in Mexico are reported in Table 4. Panel A presents the results from the DID specification (5), while Panel B shows the results for the event-study specification in (6). We report results for three dependent variables: A dummy if there was at least one Chinese investment project in the jurisdiction in column (1), the actual number of projects in column (2), and the i.h.s. investments in columns (3). In addition to location and region-by-year fixed effects, all specifications include state-specific linear trends to account for systematic patterns in the investments at the state level. We cluster our standard errors at the location level to allow for serial correlation in the unobservables within geographical units.

As reported in Panel A of Table 4, we find suggestive evidence that the ZLFN program led to a reduction in the Chinese FDI in border municipalities, with large and statistically significant negative effects for all the outcomes considered, including the inverse hyperbolic sine investments. Panel B indicates that these effects materialize in the latest years of the sample, again underscoring the delayed response of foreign investment to policy changes, and the fact that, even with the minimum wage differential began in 2019, the ZLFN minimum wage may not have been binding until the latter years. Importantly, we find no evidence

of pre-trends in any of the outcomes, which supports the parallel trends assumption behind our estimates. Overall, these findings highlight the role of place-based policies in influencing the geographic distribution of foreign direct investment, with patterns consistent with our Hypothesis 3.

6 Conclusion and Policy Implications

This study provides comprehensive evidence that the US-China Trade War significantly influenced Chinese foreign direct investment patterns in Mexico, with effects operating through sectoral reallocation toward tariff-affected products. Our findings demonstrate that Chinese firms shifted production to Mexico as an export platform for maintaining access to the US market while reducing their exposure to geopolitical tensions, with investment effects materializing several years after initial policy implementation.

The analysis reveals several important insights for both theoretical understanding and policy design. Theoretically, our results provide strong support for models of nearshoring during trade disputes while highlighting the importance of temporal dynamics in investment responses. The 3-5 year lag between tariff implementation and peak investment effects underscores the need for long-term perspectives in evaluating trade policy impacts.

From a policy perspective, our findings illustrate both the opportunities and challenges that nearshoring creates for host countries. Mexico's success in attracting Chinese investment demonstrates the potential benefits of maintaining open investment policies and leveraging comparative advantages and trade agreements. However, our separate analysis of the ZLFN program reveals how domestic policy choices can influence the geographic distribution of foreign investment within countries, showing that certain place-based policies may inadvertently reduce export-oriented investments even when broader economic conditions favor nearshoring.

Our research also contributes to broader debates about the effectiveness of trade policy in

an era of globally integrated production networks. The evidence that Chinese firms reacted to US tariffs by increasing production in Mexico suggests that unilateral trade policies may have limited effectiveness when alternative production locations are readily available. This finding has important implications for ongoing discussions about supply chain resilience, friend-shoring, and the use of economic tools for geopolitical purposes.

The geographic patterns we document have also important implications for regional development in Mexico and for policymakers in other emerging economies attracting foreign investment. While Chinese FDI inflows can generate employment opportunities, technology transfer, and integration into global value chains, our findings highlight that domestic policy can shape who benefits and where. In the case of the ZLFN program, well-intentioned measures to boost local wages inadvertently reduced export-oriented investment in affected municipalities, shifting potential job creation to other regions. More broadly, our results suggest that place-based policies can act as either complements or constraints to global investment diversion, influencing not only the volume but also the spatial distribution of the gains from nearshoring.

Several avenues for future research emerge from our analysis. First, extending the temporal scope of analysis as more post-trade war data becomes available will provide insights into the durability of nearshoring effects and potential changes in investment patterns as US-China tensions continue to evolve. Second, additional analysis of other host countries could reveal whether the patterns we identify in Mexico generalize to other nearshoring destinations. Finally, firm-level analysis could provide deeper insights into the decision-making processes that drive investment location choices during trade disputes.

As US-China tensions continue to grow and other countries grapple with similar challenges of managing great power competition, the lessons from Mexico's experience with Chinese nearshoring will become increasingly relevant for policymakers worldwide. Our study provides an empirical foundation for understanding these dynamics and designing policies that can effectively navigate the complex intersection of trade policy, investment flows, and geopolitical competition.

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Tables

Table 1: Chinese FDI and changes in tariffs: Project-level difference-in-differences and event-study specifications

	(1)	(2)	(3)	(4)	(5)
VARIABLES	Investments	Investments	Investments	Investments	Investments
A. Difference-in-diffe	rences result	S			
$Post_t \times \Delta Tariff_p$	0.052	0.052	0.048	0.033	0.036
	(0.037)	(0.037)	(0.036)	(0.030)	(0.034)
Observations	157	157	157	157	157
R-squared	0.158	0.158	0.174	0.194	0.234
B. Event-study result	ts				
$\leq 2014 \times \Delta Tariff_p$	-0.008	-0.008	0.000	0.010	0.027
•	(0.044)	(0.045)	(0.040)	(0.033)	(0.036)
2015 – $2016 \times \Delta Tariff_p$	0.007	0.007	0.009	0.016	0.002
•	(0.027)	(0.028)	(0.030)	(0.030)	(0.028)
$2019-2020 \times \Delta Tariff_p$	0.046	0.045	0.047	0.046	0.050
•	(0.036)	(0.036)	(0.036)	(0.036)	(0.037)
$2021-2022 \times \Delta Tariff_p$	0.027	0.027	0.025	0.009	0.014
-	(0.056)	(0.057)	(0.061)	(0.064)	(0.066)
2023 – $2024 \times \Delta Tariff_p$	0.064**	0.064**	0.061**	0.047**	0.050**
-	(0.028)	(0.027)	(0.026)	(0.022)	(0.025)
Observations	157	157	157	157	157
R-squared	0.163	0.163	0.178	0.199	0.240
Year FE	YES	YES	YES	YES	YES
Manufacturing FE	NO	YES	YES	YES	YES
Region FE	NO	NO	YES	YES	YES
Manufacturing trends	NO	NO	NO	YES	YES
Region trends	NO	NO	NO	NO	YES

Note: The dependent variable is the amount invested in millions of US dollars by Chinese firms in Mexico (in logs). Panel A has the results for the difference-in-differences specification in (1). Panel B has the estimates for the event-study specification in (2). Columns differ in the FE and linear trends considered. Manufacturing is a dummy for firms in the manufacturing sector. Region has four categories (north, north central, central, and south). Standard errors clustered at four-digit HS level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

Table 2: Chinese FDI and changes in tariffs: Investment-level difference-in-differences and event-study specifications controlling for the Chinese retaliatory tariffs

	(1)	(2)	(3)	(4)	(5)	
VARIABLES	Investments	Investments	Investments	Investments	Investments	
A. Difference-in-differences results						
	0.049	0.049	0.000	0.000	0.000	
$Post_t \times \Delta Tariff_p$	0.042 (0.031)	0.042 (0.031)	0.038 (0.030)	0.022	0.023	
	(0.031)	(0.031)	(0.030)	(0.025)	(0.029)	
Observations	157	157	157	157	157	
R-squared	0.170	0.170	0.186	0.208	0.250	
-						
B. Event-study result	ts					
$\leq 2014 \times \Delta Tariff_p$	0.016	0.015	0.023	0.035	0.056**	
	(0.031)	(0.031)	(0.028)	(0.025)	(0.028)	
$2015-2016 \times \Delta Tariff_p$	0.030	$0.033^{'}$	0.034	0.048	0.043	
	(0.051)	(0.051)	(0.054)	(0.054)	(0.047)	
$2019-2020 \times \Delta Tariff_p$	0.051	0.048	0.051	0.049	0.052	
	(0.049)	(0.048)	(0.049)	(0.049)	(0.050)	
2021 – $2022 \times \Delta Tariff_p$	0.023	0.024	0.024	0.006	0.012	
	(0.048)	(0.049)	(0.055)	(0.058)	(0.056)	
2023 – $2024 \times \Delta Tariff_p$	0.063**	0.062**	0.059**	0.043*	0.046*	
	(0.028)	(0.028)	(0.027)	(0.022)	(0.026)	
Observations	157	157	157	157	157	
R-squared	0.201	0.202	0.217	0.243	0.298	
Year FE	YES	YES	YES	YES	YES	
Manufacturing FE	NO	YES	YES	YES	YES	
Region FE	NO	NO	YES	YES	YES	
Manufacturing trends	NO	NO	NO	YES	YES	
Region trends	NO	NO	NO	NO	YES	

Note: The dependent variable is the amount invested in millions of US dollars by Chinese firms in Mexico (in logs). Panel A has the results for the difference-in-differences specification in (1). Panel B has the estimates for the event-study specification in (2). Columns differ in the FE and linear trends considered. All columns include as controls the Chinese retaliatory tariffs and its interaction with the post or time dummies (not reported). Manufacturing is a dummy for firms in the manufacturing sector. Region has four categories (north, north central, central, and south). Standard errors clustered at four-digit HS level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

Table 3: Chinese FDI and changes in tariffs: Heading-level difference-in-differences and event-study specifications controlling for Chinese retaliatory tariffs

VARIABLES	(1) Projects>0	(2) Projects	(3) i.h.s. inv.
			1.11.5. 1111.
A. Difference-in-diffe	rences result	ts	
$Post_t \times \Delta Tariff_h$	0.003	0.005	0.017
$1 03t_t \land \Delta 1 art f f_h$	(0.002)	(0.006)	(0.013)
	(0.002)	(0.000)	(0.010)
Observations	1,344	1,344	1,344
# headings	55	55	55
R-squared	0.193	0.256	0.224
B. Event-study result	ts		
$\leq 2014 \times \Delta Tariff_h$	0.001	-0.001	0.002
	(0.004)	(0.013)	(0.020)
2015 – $2016 \times \Delta Tariff_h$	0.001	-0.011	-0.004
	(0.003)	(0.009)	(0.014)
$2019–2020 \times \Delta Tariff_h$	-0.003	-0.008	-0.014
	(0.005)	(0.009)	(0.023)
2021 – $2022 \times \Delta Tariff_h$	0.003	0.002	0.013
	(0.004)	(0.011)	(0.025)
2023 – $2024 \times \Delta Tariff_h$	0.012*	0.015*	0.054*
	(0.006)	(0.008)	(0.030)
Observations	1,344	1,344	1,344
# headings	55	55	55
R-squared	0.209	0.286	0.243
_			
Year FE	YES	YES	YES
Heading FE	YES	YES	YES
Retaliatory tariffs	YES	YES	YES

Note: The dependent variables are a dummy if there is at least one investment by Chinese firms in Mexico in column (1), the number of investment projects in column (2), and the amount invested transformed with the inverse hyperbolic sine (i.h.s.) function in (3). Panel A has the results for the difference-in-differences specification in (3). Panel B has the estimates for the event-study specification in (4). All columns include as controls the Chinese retaliatory tariffs and its interaction with the post or time dummies (not reported). Standard errors clustered at four-digit HS level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

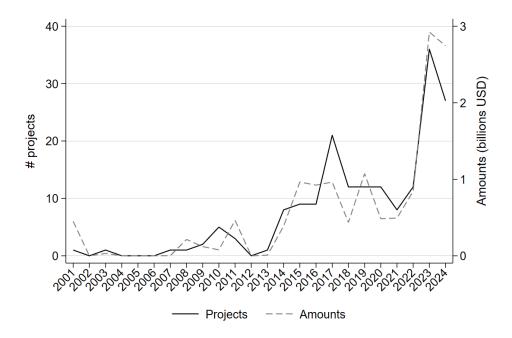
Table 4: Chinese FDI in border municipalities

	(1)	(2)	(3)				
VARIABLES	Projects>0	Projects	i.h.s. inv.				
A. Difference-in-differences results							
$Post_t \times Border_l$	-0.270*	-1.129*					
	(0.148)	(0.615)	(0.947)				
Ob	000	000	000				
Observations	888	888	888				
R-squared	0.447	0.539	0.478				
B. Event-study res	B. Event-study results						
$< 2014 \times Border_l$	0.045	0.307	0.401				
$\leq 2014 \times D01461$	(0.117)	(0.239)	(0.592)				
$2019-2021 \times Border_{l}$	-0.250	-0.295	-1.282				
2019 2021 × Dorwer	(0.179)	(0.339)	(0.967)				
$2022-2024 \times Border_{l}$	-0.228	-1.577*	-1.637				
	(0.143)	(0.924)	(1.036)				
	,	(/	,				
Observations	888	888	888				
R-squared	0.448	0.560	0.479				
Year-by-region FE	YES	YES	YES				
State-by-border FE	YES	YES	YES				
State trends	YES	YES	YES				

Note: The dependent variables are a dummy if there is at least one investment by Chinese firms in Mexico in column (1), the number of investment projects in column (2), and the inverse hyperbolic sine (i.h.s.) investment in (3). Panel A has the results for the difference-indifferences specification in (5). Panel B has the estimates for the event-study specification in (6). Standard errors clustered at state-border level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

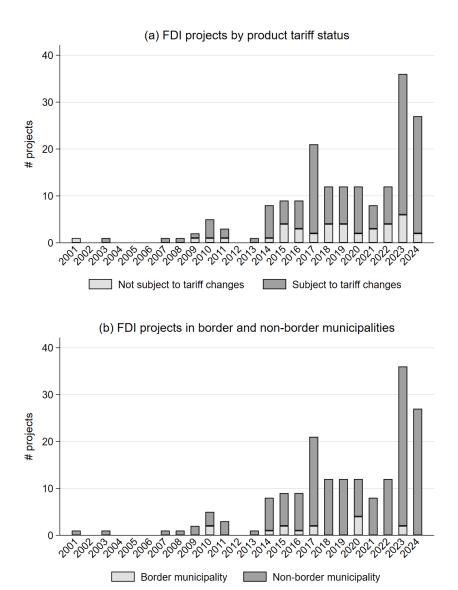
Figures

Figure 1: Number of projects and amount invested by Chinese investors in Mexico



Note: This figure plots the evolution of the number of projects (left axis) and amount invested (right axis) by Chinese firms in Mexico during 2001-2024.

Figure 2: Evolution of the number of Chinese projects exposed and not to the US-China Trade War and to the ZLFN program



Note: This figure plots the evolution of the number of projects by Chinese firms in Mexico between 2001 and 2024. In Panel (a), we differentiate between projects that experienced an increase in US tariffs to Chinese goods. In Panel (b), we split the project between those inside and outside the ZNLF region.

Figure 3: Chinese investment projects by state in Mexico



Figure 4: Chinese investment projects by state in Mexico before and after 2017



Appendix

Table A.1: Chinese FDI and exposure to US-China Trade War: Project-level difference-indifferences and event-study specifications

	(1)	(2)	(3)	(4)	(5)	
VARIABLES	Investments	Investments	Investments	Investments	Investments	
A. Difference-in-differences results						
$Post_t \times Exposed_p$	1.293	1.293	1.172	0.772	0.842	
	(0.975)	(0.976)	(0.933)	(0.840)	(0.961)	
Observations	157	157	157	157	157	
R-squared	0.149	0.149	0.164	0.184	0.224	
•						
B. Event-study resul	\mathbf{ts}					
$\leq 2014 \times Exposed_p$	0.188	0.162	0.376	0.632	1.154	
r	(1.228)	(1.238)	(1.095)	(0.867)	(0.958)	
2015 – $2016 \times Exposed_p$	0.033	0.009	0.075	0.170	-0.137	
•	(0.607)	(0.639)	(0.673)	(0.697)	(0.660)	
$20192020 \times Exposed_p$	0.270	0.201	0.203	0.186	0.274	
	(1.025)	(1.009)	(1.063)	(1.067)	(1.130)	
2021 – $2022 \times Exposed_p$	1.428	1.454	1.461	1.056	1.099	
	(1.361)	(1.385)	(1.492)	(1.614)	(1.612)	
2023 – $2024 \times Exposed_p$	1.887***	1.877***	1.786***	1.425***	1.573***	
	(0.559)	(0.560)	(0.551)	(0.497)	(0.547)	
Observations	157	157	157	157	157	
R-squared	0.164	0.164	0.178	0.194	0.242	
Year FE	YES	YES	YES	YES	YES	
Manufacturing FE	NO	YES	YES	YES	YES	
Region FE	NO	NO	YES	YES	YES	
Manufacturing trends	NO	NO	NO	YES	YES	
Region trends	NO	NO	NO	NO	YES	
1663011 0101103	110	110	110	110	110	

Note: The dependent variable is the amount invested in millions of US dollars by Chinese firms in Mexico (in logs). Panel A has the results for the difference-in-differences specification in (1). Panel B has the estimates for the event-study specification in (2) replacing in both cases the change in tariffs by a dummy equal to one if the product experienced any increase in tariffs. Columns differ in the FE and linear trends considered. Manufacturing is a dummy for firms in the manufacturing sector. Region has four categories (north, north central, central, and south). Standard errors clustered at four-digit HS level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

Table A.2: Chinese FDI and simple average of tariffs: Project-level difference-in-differences and event-study specifications

VARIABLES	(1) Investments	(2) Investments	(3) Investments	(4) Investments	(5) Investments	
A. Difference-in-differences results						
$Post_t \times \Delta Tariff_p$	0.049 (0.039)	0.049 (0.039)	$0.045 \\ (0.037)$	0.031 (0.030)	0.034 (0.034)	
Observations R-squared	$157 \\ 0.154$	$157 \\ 0.154$	157 0.170	$157 \\ 0.192$	$157 \\ 0.232$	
B. Event-study resul	ts					
$\leq 2014 \times \Delta Tariff_p$	-0.013 (0.043)	-0.013 (0.043)	-0.005 (0.040)	0.004 (0.033)	0.020 (0.036)	
$20152016 \times \Delta Tariff_p$	0.043) 0.002 (0.027)	0.043) 0.002 (0.028)	0.040) 0.005 (0.029)	0.011 (0.029)	(0.030) -0.003 (0.027)	
2019–2020 × $\Delta Tariff_p$	0.034 (0.037)	0.033 (0.037)	0.035 (0.036)	0.034 (0.036)	0.038 (0.038)	
$20212022 \times \Delta Tariff_p$	0.023 (0.059)	0.023 (0.060)	0.022 (0.064)	0.006 (0.067)	0.011 (0.068)	
$2023 – 2024 \times \Delta Tariff_p$	0.061* (0.031)	0.060* (0.031)	0.057* (0.029)	0.044* (0.024)	0.049* (0.026)	
Observations R-squared	$157 \\ 0.160$	157 0.160	157 0.175	157 0.196	157 0.238	
Year FE Manufacturing FE Region FE Manufacturing trends	YES NO NO NO	YES YES NO NO	YES YES YES NO	YES YES YES YES	YES YES YES YES	
Region trends	NO	NO NO	NO NO	NO	YES	

Note: The dependent variable is the amount invested in millions of US dollars by Chinese firms in Mexico (in logs). Panel A has the results for the difference-in-differences specification in (1). Panel B has the estimates for the event-study specification in (2). The change in tariffs at the 6-digit HS level corresponds to the simple average of the tariffs at the 10-digit level. Columns differ in the FE and linear trends considered. Manufacturing is a dummy for firms in the manufacturing sector. Region has four categories (north, north central, central, and south). Standard errors clustered at four-digit HS level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.

Table A.3: Chinese FDI and simple average of tariffs: Heading-level difference-in-differences and event-study specifications controlling for Chinese retaliatory tariffs

VARIABLES	(1) Projects>0	(2) Projects	(3) i.h.s. inv.			
A. Difference-in-differences results						
A. Difference-in-difference-in	rences result	ts				
$Post_t \times \Delta Tariff_h$	0.002	0.003	0.013			
Toott N AT arrej jn	(0.003)	(0.007)	(0.015)			
	(3.333)	(3.33.)	(0.0_0)			
Observations	1,344	1,344	1,344			
# headings	55	55	55			
R-squared	0.191	0.261	0.222			
B. Event-study result	ts					
< 2014 × A.T f.f.	0.000	0.000	0.002			
$\leq 2014 \times \Delta Tariff_h$	$0.000 \\ (0.005)$	0.002 (0.014)	0.003 (0.022)			
2015 – $2016 \times \Delta Tariff_h$	-0.002	-0.014)	-0.014			
$2019-2010 \times \Delta I \text{ at if } f_h$	-0.002 (0.002)	(0.013)	-0.014 (0.014)			
$2019-2020 \times \Delta Tariff_h$	-0.002	-0.012	-0.029			
$2019 \ 2020 \times \Delta 1 \ arr ijjh$	(0.005)	(0.009)	(0.026)			
$2021-2022 \times \Delta Tariff_h$	0.002	0.004	0.017			
	(0.004)	(0.012)	(0.029)			
$2023-2024 \times \Delta Tariff_h$	0.011*	0.015*	0.053*			
<i>y y</i> 10	(0.006)	(0.009)	(0.030)			
	, ,	, ,	,			
Observations	1,344	1,344	1,344			
# headings	55	55	55			
R-squared	0.211	0.317	0.245			
X DD	NEC.	T C	T.T.C			
Year FE	YES	YES	YES			
Heading FE	YES	YES	YES			
Retaliatory tariffs	YES	YES	YES			

Note: The dependent variables are a dummy if there is at least one investment by Chinese firms in Mexico in column (1), the number of investment projects in column (2), and the amount invested transformed with the inverse hyperbolic sine (i.h.s.) function in (3). Panel A has the results for the difference-in-differences specification in (3). Panel B has the estimates for the event-study specification in (4). The change in tariffs at the heading level corresponds to the simple average of the tariffs at the 10-digit level. All columns include as controls the Chinese retaliatory tariffs and its interaction with the post or time dummies (not reported). Standard errors clustered at four-digit HS level in parentheses. *** p < 0.01, ** p < 0.05, * p < 0.1.